

# An extension of the Benjamini and Hochberg procedure using data-driven optimal weights with grouped hypothesis

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- 1 Introduction : BH and weighting
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## The well-known BH procedure

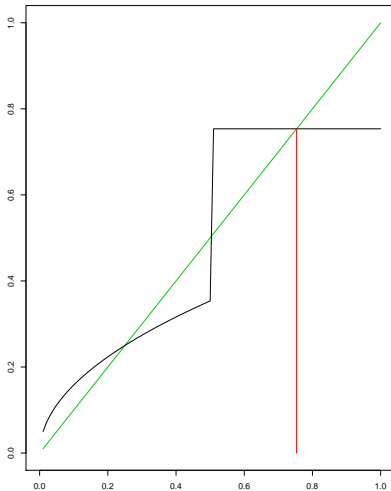
- Order p-values :  $p_{(1)} \leq \dots \leq p_{(m)}$
- Compute  $\hat{k} = \max\{k : p_{(k)} \leq \alpha k/m\}$
- Reject all  $p_i \leq \alpha \frac{\hat{k}}{m}$
- FDR control at level  $\pi_0 \alpha$  when wPRDS

### Another formulation

$\frac{\hat{k}}{m} = \max\{u : \hat{G}(u) \geq u\} := \mathcal{I}(\hat{G})$  where

$$\hat{G} : u \mapsto m^{-1} \sum_{i=1}^m \mathbb{1}_{\{p_i \leq \alpha u\}}, u \in [0, 1]$$

# An illustration of $\mathcal{I}(F)$



# Weighted-BH

With given weights  $(w_i)_{1 \leq i \leq m}$  such that  $\sum_i w_i = m$  (called a weight vector), form

$$\widehat{G}_w : u \mapsto m^{-1} \sum_{i=1}^m \mathbb{1}_{\{p_i \leq \alpha u w_i\}}$$

and reject all  $p_i \leq \alpha \hat{u} w_i$  with  $\hat{u} = \mathcal{I}(\widehat{G}_w)$ .

# Weighted-BH

A generalization : weight functions

From Roquain and Van De Wiel 2009 :

Take a function  $W$  such that  $(W_i(u))_i$  is a weight vector for all  $u$  and

$$\widehat{G}_W : u \mapsto m^{-1} \sum_{i=1}^m \mathbb{1}_{\{p_i \leq \alpha u W_i(u)\}}$$

is non-decreasing, then reject all  $p_i \leq \alpha \widehat{u} W_i(\widehat{u})$  with  $\widehat{u} = \mathcal{I}(\widehat{G}_W)$ .

## Weighted-BH

A practical way to compute  $\mathcal{I}(\hat{G}_W)$

- No need to compute  $W(u)$  for each  $u$  !

For each  $k \in \llbracket 1, m \rrbracket$ , compute the  $\frac{p_i}{W_i(\frac{k}{m})}$  and take  $q_r$  the  $r$ -th smallest. Let  $q_0 = 0$ .

Then  $\mathcal{I}(\hat{G}_W) = m^{-1} \max\{k \in \llbracket 0, m \rrbracket : q_k \leq \alpha \frac{k}{m}\}$ .



## Optimal weighting

- Unconditional model : each hypothesis is null with proba  $\pi_0$ .
- Consider the procedure  $R_m^u$  rejecting  $p_i$  if  $p_i \leq \alpha u w_i$  for all  $u$ .
- Its power is  $\text{Pow}_w(u) := (1 - \pi_0) m^{-1} \sum_i^m F_i(\alpha u w_i)$  ( $F_i$  the c.d.f. under the alternative).
- Maximize it for all  $u$  :

Definition of optimal weights :

$$W^*(u) = \underset{(w_i) \text{ s.t. } \sum_i^m w_i = m}{\text{argmax}} \text{Pow}_w(u)$$

# Optimal weighting

## Existence and uniqueness

### Theorem (Roquain and Van De Wiel 2009)

Assume the following :

- $F_i$  is strictly concave and continuous on  $[0, 1]$
- $F_i$  has a derivative  $f_i$  on  $(0, 1)$
- $f_i(0^+)$  is constant for all  $i$ , same for  $f_i(1^-)$
- $\lim_{y \rightarrow f_i(0^+)} \frac{f_j^{-1}(y)}{f_i^{-1}(y)}$  exists in  $[0, \infty]$  for all  $i, j$

Then we have existence, uniqueness and continuity of  $W^*$ , and  $u \mapsto uW_i^*(u)$  is non-decreasing.

# Optimal weighting

## Existence and uniqueness

### Proof ideas

Compute an explicit formula using the Lagrange multiplier method :

$$L(\lambda, w) = m^{-1} \sum_{i=1}^m F_i(\alpha u w_i) - \lambda \left( \sum_{i=1}^m w_i - m \right)$$

gives us

$$W_i^*(u) = \frac{1}{\alpha u} f_i^{-1}(\Psi^{-1}(\alpha u))$$

where  $\Psi(x) = m^{-1} \sum_{i=1}^m f_i^{-1}(x)$ .

# Optimal weighting

The main problem and the resulting motivation

- The distribution under the alternative  $F_i$  needs to be known to compute the  $W^*$ .
- Goal : estimate  $W^*$  without the knowledge of the alternative and obtain asymptotical results on FDR control and power for the associated weighted-BH procedure.
- Leads to data-driven optimal weighting.

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## Data-driven optimal weighting

- Assume that the p-values have uniform distribution under the null.

Main idea :

$W^*(u)$  is also the unique maximizer of

$$G_w(u) = \mathbb{E} \left[ \widehat{G}_w(u) \right] = \pi_0 m^{-1} \sum_i^m \max(\alpha u w_i, 1) + \text{Pow}_w(u)$$

the mean proportion of rejections done by the procedure  $R_m^u$ .

So we can estimate  $W^*$  by maximizing  $G_w$ 's empiric counterpart  $\widehat{G}_w$ .

# Data-driven optimal weighting

## Grouping hypotheses

### Key assumption

$G$  groups of sizes  $(m_g)_{1 \leq g \leq G}$ , where p-values have the same distribution.

Examples :

- The Adequate Yearly Progress data set where grouping schools by size avoids a preference for large schools.
- Search for differently expressed genes between individuals with normal copy number or amplified one. Tests are more efficient for genes with ratio normal vs amplified copy numbers near 1.

# Data-driven optimal weighting

## More technical hypotheses

To obtain asymptotic results on  $\widehat{W}^*$ , we assume the following :

- p-values are independent.
- The previous hypothesis made by Roquain and Van De Wiel remain, and in addition  $f_g(0^+) = \infty \forall g$ .
- $\frac{m_g}{m} \xrightarrow{m \rightarrow \infty} \pi_g > 0$ .

All the following proofs inspired by Roquain and Van De Wiel 2009, Zhao and Zhang 2014 and Hu, Zhao, and Zhou 2010.



## The two main results

### Theorem (FDR control)

The False Discovery Proportion converges to  $\pi_0\alpha$  almost surely so the FDR too by dominated convergence.

### Theorem (optimal power)

Note by  $\text{Pow}(W)$  the power of a BH procedure using a weight function  $W$ . Note by  $\mathscr{W}$  the set of all sequences  $(w^{(m)})$  such that  $\sum m_g w_g^{(m)} = m$ . Then :

$$\lim_{m \rightarrow \infty} \text{Pow}(\widehat{W}^*) \geq \sup_{(w^{(m)}) \in \mathscr{W}} \limsup_{m \rightarrow \infty} \text{Pow}(w^{(m)}).$$

## Some notations

- From now  $W^*$  is the asymptotic optimal weight when the  $F_g$  are known :

$$\begin{aligned}W^*(u) &= \operatorname{argmax}_{w: \sum \pi_g w_g = 1} G_w^\infty(u) \\ &= \operatorname{argmax}_{w: \sum \pi_g w_g = 1} \sum_g \pi_g D_g(\alpha u w_g)\end{aligned}$$

with  $D_g(\cdot) = \pi_0 \max(\cdot, 1) + (1 - \pi_0)F_g(\cdot)$ .

- $P_W^\infty(u) = (1 - \pi_0) \sum_g \pi_g F_g(\alpha u W_g(u))$ .
- $\hat{u} = \mathcal{I}(\hat{G}_{\hat{W}^*})$  and  $u^* = \mathcal{I}(G_{W^*}^\infty)$ .

# A chain of technical results

## A first lemma

$$\sup_{u \in [0,1]} \sup_{w \in (\mathbb{R}^+)^G} \left| \widehat{G}_w(u) - G_w^\infty(u) \right| \xrightarrow{a.s.} 0$$

by Glivenko-Cantelli theorem and  $\frac{m_g}{m} \rightarrow \pi_g$ .

# The main technical proposition

## Proposition

$$\sup_{u \in [0,1]} \left| \widehat{G}_{\widehat{W}^*}(u) - G_{W^*}^\infty(u) \right| \xrightarrow{a.s.} 0$$

or, equivalently,

$$\sup_{u \in [0,1]} \left| G_{\widehat{W}^*}^\infty(u) - G_{W^*}^\infty(u) \right| \xrightarrow{a.s.} 0.$$

# The main technical proposition

## Proof ideas

- Play with the triangular inequality and remove the absolute values when able by using the maximality of  $\widehat{G}_{\widehat{W}^*}(u)$  and  $G_{W^*}^\infty(u)$

### Problem

They are not maxima on the same sets :

$$K^m = \{w : m^{-1} \sum m_g w_g = 1\} \text{ versus } K^\infty = \{w : \sum \pi_g w_g = 1\}$$

# The main technical proposition

## Proof ideas

- We introduce two shifts  $\delta(\mathbf{u}) = \sum \pi_g \widehat{W}_g^*(\mathbf{u}) - 1$  and  $\delta'(\mathbf{u}) = \sum \frac{m_g}{m} W_g^*(\mathbf{u}) - 1$ .
- Then we form shifted weights  $\widehat{W}^\sim(\mathbf{u}) = \widehat{W}^*(\mathbf{u}) - \delta(\mathbf{u}) \in K^\infty$  and  $W^\sim(\mathbf{u}) = W^*(\mathbf{u}) - \delta'(\mathbf{u}) \in K^m$ .

# The main technical proposition

## Final ideas

- Make appear  $\left| G_{\widehat{W}^\sim}^\infty(u) - G_{\widehat{W}^*}^\infty(u) \right| = G_{\widehat{W}^*}^\infty(u) - G_{\widehat{W}^\sim}^\infty(u)$ .
- End up with  $\sup_u \left| G_{\widehat{W}^\sim}^\infty(u) - G_{\widehat{W}^*}^\infty(u) \right| \leq$   
 $\sup_u \left( \widehat{G}_{\widehat{W}^\sim}(u) - \widehat{G}_{\widehat{W}^*}(u) \right) + o_{a.s.}(1)$ .
- Use that  $\widehat{G}_{\widehat{W}^\sim}(u) - \widehat{G}_{\widehat{W}^*}(u) \leq 0$ .  $\square$

## The second important proposition

### Proposition

$$\hat{u} \xrightarrow[m \rightarrow \infty]{a.s.} u^*$$

from which we deduce  $\widehat{G}_{\widehat{W}^*}(\hat{u}) \xrightarrow{a.s.} G_{W^*}^\infty(u^*)$  by continuity.

Note  $X_m = \sup_{u \in [0,1]} \left| \widehat{G}_{\widehat{W}^*}(u) - G_{W^*}^\infty(u) \right| \xrightarrow{a.s.} 0$ , take a  $\delta$  in  $(0, u^*)$ ,  
note  $u^0 = u^* - \delta$  and for all  $\delta' \geq \delta$ ,  $u' = u^* + \delta'$ .



# The second important proposition

## Proof

- $s_\delta = \max_{\delta' \geq \delta} (G_{W^*}^\infty(u') - u') < 0$  because if  $s_\delta = 0$  it would contradict  $u^*$  maximality.
- $\sup_{\delta' \geq \delta} (\widehat{G}_{\widehat{W}^*}(u') - u') \leq s_\delta + X_m \rightarrow s_\delta < 0$
- So when  $m \rightarrow \infty$  we must have  $\hat{u} < u^* + \delta$ .

# The second important proposition

## Proof

- $G_{W^*}^\infty(u^0) \geq G_w^\infty(u^0)$  with  $w = W^*(u^*)$  by maximality.
- $G_w^\infty(u^0) = \frac{G_w^\infty(u^0)}{u^0} u^0 > \frac{G_w^\infty(u^*)}{u^*} u^0 = u^0$  by strict concavity.
- $\widehat{G}_{\widehat{W}^*}(u^0) - u^0 \geq G_{W^*}^\infty(u^0) - u^0 - X_m \rightarrow G_{W^*}^\infty(u^0) - u^0 > 0$ .
- So when  $m \rightarrow \infty$  we must have  $\hat{u} > u^* - \delta$ .  $\square$

## Third and last proposition

We have shown that  $\widehat{G}_{\widehat{W}^*}(\hat{u}) \xrightarrow{a.s.} u^*$ , that is for the denominator of the FDP. Showing that the numerator converges to  $\pi_0 \alpha u^*$  is straightforward after this :

### Proposition

$$\widehat{W}^*(\hat{u}) \xrightarrow{a.s.} W^*(u^*),$$

or, equivalently,

$$\widehat{W}^{\sim}(\hat{u}) \xrightarrow{a.s.} W^*(u^*).$$

## Third and last proposition

### Proof ideas

- One can show with the previous results and the triangular inequality that  $\left| G_{\widehat{W}^{\sim}(\hat{u})}^{\infty}(u^*) - G_{W^*}^{\infty}(u^*) \right| \xrightarrow{a.s.} 0$ .
- By contradiction, if  $\widehat{W}^{\sim}(\hat{u}) \xrightarrow{a.s.} W^*(u^*)$  then we find a  $w^l \neq W^*(u^*)$  maximizing  $G_w^{\infty}(u^*)$  but  $W^*(u^*)$  is unique.  $\square$

# Optimality in power

## Proof ideas

- First,  $\text{Pow}(\widehat{W}^*) = \mathbb{E} \left[ \widehat{P}_{\widehat{W}^*}(\hat{u}) \right]$  where  $\widehat{P}_W(u)$  is  $m^{-1}$  times the number of true alternative rejected.
- $\widehat{P}_{\widehat{W}^*}(\hat{u}) \xrightarrow{a.s.} P_{W^*}^\infty(u^*)$ .
- For each accumulation point for  $\text{Pow}(w^{(m)})$  there is an accumulation point  $w$  for  $w^{(m)}$ .
- $\hat{u}^{(m'')} \xrightarrow{a.s.} \mathcal{I}(G_w^\infty)$  and then
- $\widehat{P}_{w^{(m'')}}(\hat{u}^{(m'')}) \xrightarrow{a.s.} P_w^\infty(\mathcal{I}(G_w^\infty)) \leq P_{W^*}^\infty(\mathcal{I}(G_w^\infty)) \leq P_{W^*}^\infty(u^*)$ .  $\square$

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## Computation

- Fix a  $u$ , form  $\tilde{p}_{gi} = \frac{p_{gi}}{\alpha u}$  and order the  $\tilde{p}_{gi}$  in each group :

$$\tilde{p}_{g,1} \leq \dots \leq \tilde{p}_{g,m_g}.$$

Also note  $\tilde{p}_{g,0} = 0$ .

- Maximize over  $w : \sum m_g w_g = m \iff$  maximize over  $w : \sum m_g w_g \leq m$ .
- If  $\forall g, \tilde{p}_{g,1} > m$ , stop and no rejection. If  $\exists g, \tilde{p}_{g,1} \leq m$ , continue and at least one rejection.

# Computation

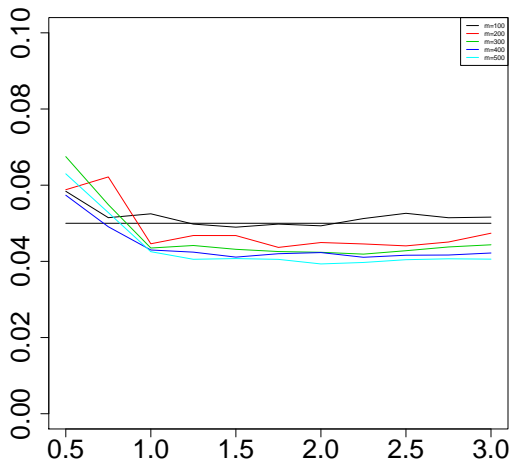
- Form all G-tuples  $\mathbf{j} : \sum j_g = 2$  and check if there is one  $\mathbf{j}$  such that  $\sum m_g \tilde{p}_{g,j_g} \leq m$ 
  - If there is one, at least 2 rejections and continue with G-tuples of sum equal to 3.
  - If not, 1 rejection and use a  $w_g = \tilde{p}_{g,j_g}$  with a  
$$\mathbf{j} = (0, \dots, 0, \overbrace{1}^{h\text{-th position}}, 0, \dots, 0)$$
 such that  $\tilde{p}_{h,1} \leq m$ .
- Reminder : the only values of  $u$  that need to be computed are  $1, \frac{m-1}{m}, \dots, \frac{1}{m}$ .



# First simulations : $\alpha = 0.05$ , 70% null hypothesis

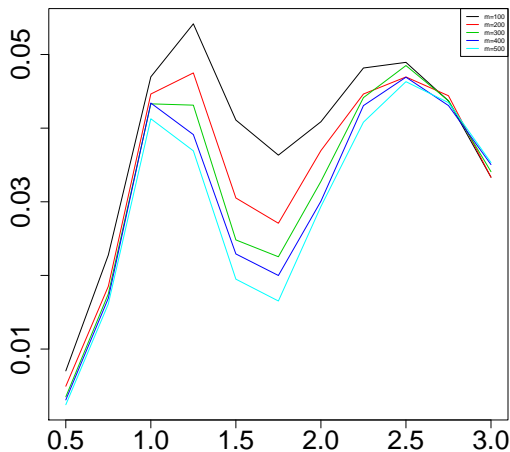
FDR plot, 2 groups,  $\pi_1 = \pi_2 = 0.5$

- x axis : the  $\bar{\mu}$  parameter.  $\mu_1 = \bar{\mu}$  and  $\mu_2 = 2\bar{\mu}$ .
- y axis : the FDR.



# First simulations : $\alpha = 0.05$ , 70% null hypothesis

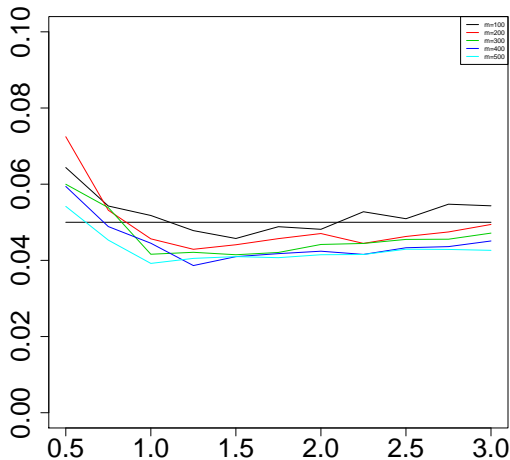
Relative power plot, 2 groups,  $\pi_1 = \pi_2 = 0.5$



- x axis : the  $\bar{\mu}$  parameter.  $\mu_1 = \bar{\mu}$  and  $\mu_2 = 2\bar{\mu}$ .
- y axis : the difference in power between our procedure and the BH procedure.

# First simulations : $\alpha = 0.05$ , 70% null hypothesis

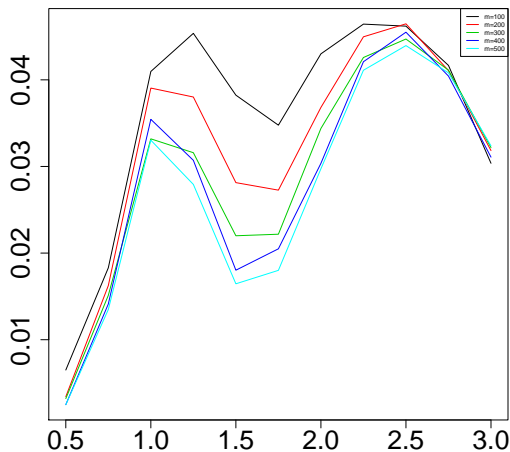
FDR plot, 2 groups,  $\pi_1 = 0.3, \pi_2 = 0.7$



- x axis : the  $\bar{\mu}$  parameter.  $\mu_1 = \bar{\mu}$  and  $\mu_2 = 2\bar{\mu}$ .
- y axis : the FDR.

## First simulations : $\alpha = 0.05$ , 70% null hypothesis

Relative power plot, 2 groups,  $\pi_1 = 0.3, \pi_2 = 0.7$



- x axis : the  $\bar{\mu}$  parameter.  $\mu_1 = \bar{\mu}$  and  $\mu_2 = 2\bar{\mu}$ .
- y axis : the difference in power between our procedure and the BH procedure.

## Some perspectives

- Estimate  $\pi_0$  to control the FDR at level  $\alpha$  instead of  $\alpha\pi_0$ .
- A different  $\pi_0$  in each group.
- Use wPRDS instead of independence.
- Optimize the computation.

The end

Thank you for your attention !

## Bibliography

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